Estimating Average and Local Average Treatment Effects of Education When Compulsory Schooling Laws Really Matter: Corrigendum

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Philip Oreopoulos Department of Economics, University of British Columbia <u>philip.oreopoulos@ubc.ca</u>

In the March 2006 edition of this journal, I published an article entitled "Estimating Average and Local Average Treatment Effects of Education when Compulsory School Laws Really Matter" (American Economic Review, Vol. 96, No. 1, March 2006, pp. 152-175). In the last year, some colleagues have informed me that they have had difficulty replicating the UK results using the code I provided in a data appendix.¹ Through these discussions I learned that a few sampling restrictions that were mentioned in the paper were not in the code, and that some datasets were not merged correctly (for example, individuals were matched based on person and household identifiers, but not family identifiers). The British earnings measure for 1994 was also accidentally dropped. This corrigendum therefore updates the code for producing a revised set of UK results which are qualitatively similar to the original results. The revised output does not affect the discussion or conclusions of the original article.

One of the primary ideas behind the original article is that the remarkably large response from changes to compulsory schooling laws in the UK provides a rare

¹ I thank Paul Devereux, Heather Royer, Joseph Shapiro, and Raymond Guiteras for pointing me to these mistakes. I am a strong supporter of making available code for replication purposes, and I am grateful that these errors were identified using the paper's data appendix. Part of the difficulty reconciling the results was due to keeping only an aggregated version of the data. For the revised results here, I include the full micro dataset.

opportunity to measure average returns to schooling for a more general population than compared to previous papers using instrumental variables methodology. First stage estimates suggest that raising the school leaving age from 14 to 15 in 1947 for Great Britain and in 1957 for Northern Ireland affected between 40 to 50 percent of the general cohort population. The original article concluded that the similar UK returns to compulsory schooling estimates compared to those from the U.S. and Canada (with significantly less affected by the policy changes) suggests that the average treatment effect of an additional year of high school is about as large as local average treatment effect estimated for these countries (at least for the particular birth cohorts examined in the study).

Table 1 shows the estimated first stage effects of the policy changes on the number of years of schooling, the reduced form effects of the policy changes on log earnings, and the instrumental variables estimates for the returns to compulsory schooling for Great Britain and the UK. The originally published results are reported in Panel A. Panel B reports the revised results using the full sample of individuals from the 1984 to 1998 British and Northern Irish General Household Surveys, aged 28 to 64, born between 1921 and 1951 (the same sample restrictions reported in the original paper).² The results for Northern Ireland are almost identical. The first stage effects for Great Britain are also about the same, but the returns to schooling estimates are lower – around 7 percent instead of 15 percent. The 95 percent confidence region around these estimates are quite wide, ranging from returns as low as 0 and as high as 15 percent. The revised difference-

 $^{^{2}}$ While the published paper noted that the 1983 GHHS was included for the sample, income that year was recorded as missing. So, in effect, the sample included only the 1984 to 1998 GHHS.

in-difference estimates from combining the Great Britain and Northern Ireland samples are slightly lower, but still above 10 percent, with standard errors around 0.03.

Table 2 shows the same set of estimates but with three alternative sample specifications that I believe are equally justifiable. Panel A shows results after restricting the sample to individuals that left full-time schooling by age 18 or less, with the logic that since the first stage results suggest that the policy changes affected whether individuals left full time schooling before age 15 but not before age 16 (these effects are shown in the original article and revised set of tables and figures), this alternative restriction may improve precision by dropping individuals not likely impacted by the policy shift. Panel B shows results from dropping individuals aged 61 to 64. Earnings from older workers close to retirement may be more volatile. Panel C shows results from adding more data using more recent years of the British and Northern Irish General Household Surveys that were not available when I began the study (1999 to 2006).

The resulting first stage effects for Great Britain in Panel A from dropping individuals that left school after age 18 are generally not different from the full sample in Panel B of Table 1, except the first stage effects from the school leaving age change in Northern Ireland are lower, and the instrumental variables estimates have larger standard errors. Dropping individuals over 60 years of age in Panel B generates higher point estimates for the reduced form and returns to compulsory schooling British estimates. The Northern Ireland estimates are generally unchanged. Adding the additional survey data in Panel C notably improves precision and leads to significant returns to education estimates that range from 11 to 13 percent for Great Britain and 13 to 18 percent for Northern Ireland. Finally, Panel D combines all three alternative sample specifications,

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which leads to similar point estimates compared to the original article, but large standard errors for the Northern Ireland sample.

As mentioned in the original article, the regression discontinuity approach leads to greater imprecision than the difference-in-difference approach, given that earnings are tapering off for successively older birth cohorts at the time the discontinuity occurs. Since the confidence intervals include values lower than the point estimates for Canada and the U.S., the analysis requires considering the robustness and general patterns of the results under alternative samples, methodologies, and conditions. In my opinion, the results presented here, the robustness checks presented in this article's appendix, and the results presented in Oreopoulos (2007) using the Eurobarometer Surveys, point to clear evidence of substantial returns to compulsory schooling – between 8 and 15 percent — for individuals affected by these policy changes.

I have created a new set of tables and figures from the specification above that adds additional survey years from the General Household Survey, as shown in Table 2, Panel C). These can be found in the data appendix. Code to replicate these results, along with the full micro dataset, is also provided in the data appendix. The results still support the conclusions and discussion that I drew using tables of the published version.

References

Oreopoulos, Philip. "Do Dropouts Drop Out Too Soon? Wealth, Health, and Happiness from Compulsory Schooling," Journal of Public Economics 2007, 91, (11-12), 2213-2229.

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Great Britain and Northern Ireland, Baseline Sample										
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	
	(First Stage) Dependent Variable: Age Finished Full Time School			Depende	(Reduced Form) Dependent Variable: Log Annual Earnings			(IV Returns to Compulsory Schooling) Dependent Variable: Log Annual Earnings		
				P	anel A: Published Resu	lts				
Great Britain [n=57624]	0.440 [0.065]***	0.436 [0.071]***	0.453 [0.076]***	0.065 [0.025]**	0.064 [0.026]**	0.042 [0.043]	0.147 [0.061]**	0.145 [0.063]**	0.149 [0.064]**	
Northern Ireland [n=8921]	0.397 [0.074]***	0.391 [0.073]***	0.353 [0.100]***	0.054 [0.027]*	0.074 [0.025]***	0.074 [0.045]	0.135 [0.071]*	0.187 [0.070]**	0.21 [0.135]	
G. Britain and N. Ireland with N. Ireland Fixed Effect [n=66185]	0.418 [0.040]***	0.397 [0.043]***	0.401 [0.045]***	0.073 [0.016]***	0.058 [0.016]***	0.059 [0.018]***	0.174 [0.042]***	0.149 [0.044]***	0.148 [0.046]***	
			Panel B: Revised Re	esults with Baseline Sam	ple: 1921 - 1951 Birth (Cohorts aged 28-64 in th	e 1984 - 1998 GHHS			
Great Britain [n=55088]	0.408 [0.063]***	0.408 [0.064]***	0.435 [0.073]***	0.029 [0.016]*	0.025 [0.020]	0.032 [0.021]	0.069 [0.040]*	0.066 [0.050]	0.067 [0.049]	
Northern Ireland n=[8954]	0.456 [0.104]***	0.444 [0.105]***	0.413 [0.082]***	0.059 [0.040]	0.081 [0.034]**	0.074 [0.045]	0.129 [0.076]*	0.18 [0.062]***	0.179 [0.096]*	
G. Britain and N. Ireland with N. Ireland Fixed Effect [n=64042]	0.437 [0.043]***	0.44 [0.044]***	0.451 [0.044]***	0.058 [0.013]***	0.045 [0.014]***	0.044 [0.014]***	0.132 [0.031]***	0.105 [0.030]***	0.097 [0.030]***	
Birth Cohort Polynomial Controls	Quartic	Quartic	Quartic	Quartic	Quartic	Quartic	Quartic	Quartic	Quartic	
Age Polynomial Controls	None	Quartic	None	None	Quartic	None	None	Quartic	None	
Age Dummies	No	No	Yes	No	No	Yes	No	No	Yes	

Table 1 First Stage, Reduced Form, and IV Estimates for Returns to Compulsory Schooling Great Britain and Northern Ireland, Baseline Sample

Notes: The dependent variables are age left full-time education and log annual earnings. Each regressions includes controls for a birth cohort quartic polynomial and and indicator whether a cohort faced a school leaving age of 15 at age 14. Columns (2), (3), (5), (6), (8) and (9) also include age controls: a quartic polynomial and fixed effects where indicated. Each regression includes the sample of 25 to 64 year olds from the 1984 through 1998 General Household Surveys, who were aged 14 between 1935 and 1965. Data are first aggregated into cell means and weighted by cell size. Regressions are clustered by birth cohort and region (Britian or N. Ireland). [n = sample size]

Great Britain and Northern Ireland, Alternative Sample Conditions									
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
	Dependent V	(First Stage) Variable: Age Finished Ful	ll Time School	Depende	(Reduced Form) ent Variable: Log Annual	Earnings		eturns to Compulsory Sch ent Variable: Log Annual	
			Panel A: 1921	- 1951 Birth Cohorts ag	ed 28-64 in the 1984 - 19	998 GHHS, left full time	schooling < 19		
Great Britain [n=46760]	0.437 [0.080]***	0.438 [0.079]***	0.44 [0.078]***	0.032 [0.015]**	0.029 [0.019]	0.037 [0.020]*	0.073 [0.042]*	0.062 [0.050]	0.084 [0.053]
Northern Ireland [n=7676]	0.24 [0.068]***	0.24 [0.068]***	0.274 [0.072]***	0.03 [0.046]	0.054 [0.037]	0.059 [0.048]	0.123 [0.194]	0.224 [0.155]	0.203 [0.199]
G. Britain and N. Ireland with N. Ireland Fixed Effect	0.478 [0.045]***	0.477 [0.045]***	0.48 [0.043]***	0.058 [0.013]***	0.044 [0.013]***	0.043 [0.014]***	0.123 [0.030]***	0.091 [0.031]***	0.091 [0.032]***
[n=54436]				Panel B: 1921 - 1951 Bi	rth Cohorts aged 28-60 i	in the 1984 - 1998 GHHS	5		
Great Britain [n=51643]	0.376 [0.054]***	0.376 [0.053]***	0.411 [0.063]***	0.047 [0.020]**	0.044 [0.022]*	0.052 [0.023]**	0.108 [0.057]*	0.144 [0.079]*	0.107 [0.062]*
Northern Ireland [n=8311]	0.434 [0.107]***	0.436 [0.110]***	0.372 [0.094]***	0.051 [0.044]	0.085 [0.036]**	0.083 [0.046]*	0.094 [0.097]	0.179 [0.064]**	0.226 [0.098]**
G. Britain and N. Ireland with N. Ireland Fixed Effect [n=59954]	0.411 [0.043]***	0.41 [0.042]***	0.425 [0.042]***	0.064 [0.014]***	0.046 [0.014]***	0.046 [0.014]***	0.159 [0.038]***	0.112 [0.037]***	0.11 [0.037]***
				Panel C: 1921 - 1951 Bi	rth Cohorts aged 28-64 i	in the 1984 - 2006 GHHS	8		
Great Britain [n=73954]	0.495 [0.074]***	0.457 [0.065]***	0.472 [0.069]***	0.055 [0.015]***	0.052 [0.014]***	0.056 [0.017]***	0.112 [0.034]***	0.111 [0.033]***	0.125 [0.040]***
Northern Ireland [n=8954]	0.456 [0.104]***	0.444 [0.105]***	0.413 [0.082]***	0.059 [0.040]	0.081 [0.034]**	0.074 [0.045]	0.129 [0.076]*	0.18 [0.062]***	0.179 [0.096]*
G. Britain and N. Ireland with N. Ireland Fixed Effect [n=82908]	0.491 [0.042]***	0.475 [0.044]***	0.485 [0.042]***	0.02 [0.015]	0.065 [0.013]***	0.065 [0.013]***	0.041 [0.032]	0.133 [0.027]***	0.135 [0.028]***
			Panel D: 192	l - 1951 Birth Cohorts ag	ged 28-60 in the 1979 - 2	006 GHHS, left full time	e schooling < 19		
Great Britain [n=54982]	0.428 [0.086]***	0.428 [0.086]***	0.431 [0.085]***	0.063 [0.018]***	0.047 [0.023]**	0.052 [0.024]**	0.191 [0.074]**	0.118 [0.073]	0.133 [0.076]*
Northern Ireland [n=7081]	0.199 [0.054]***	0.209 [0.056]***	0.227 [0.055]***	0.022 [0.048]	0.057 [0.036]	0.068 [0.048]	0.168 [0.204]	0.313 [0.127]**	0.311 [0.184]
G. Britain and N. Ireland with N. Ireland Fixed Effect [n=62063]	0.474 [0.047]***	0.481 [0.048]***	0.483 [0.046]***	0.015 [0.018]	0.062 [0.015]***	0.062 [0.014]***	0.035 [0.040]	0.131 [0.036]***	0.127 [0.035]***
Birth Cohort Polynomial Controls	Quartic	Quartic	Quartic	Quartic	Quartic	Quartic	Quartic	Quartic	Quartic
Age Polynomial Controls	None	Quartic	None	None	Quartic	None	None	Quartic	None
Age Dummies	No	No	Yes	No	No	Yes	No	No	Yes

Table 2
First Stage, Reduced Form, and IV Estimates for Returns to Compulsory Schooling
Great Britain and Northern Ireland, Alternative Sample Conditions

Notes: The dependent variables are age left full-time education and log annual earnings. Each regressions includes controls for a birth cohort quartic polynomial and and indicator whether a cohort faced a school leaving age of 15 at age 14. Columns (2), (3), (5), (6), (8) and (9) also include age controls: a quartic polynomial and fixed effects where indicated. Each regression includes the sample from the 1984 through 1998 General Household Surveys, who were aged 14 between 1935 and 1965. Data are first aggregated into cell means and weighted by cell size. Regressions are clustered by birth cohort and region (Britian or N. Ireland). [n = sample size]

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
		(First Stage) ariable: Age Finished Fu			(Reduced Form) ent Variable: Log Annual		Initial Sample Size
Great Britain	0.495 [0.074]***	0.457 [0.065]***	0.472 [0.069]***	0.055 [0.015]***	0.052 [0.014]***	0.056 [0.017]***	73954
Northern Ireland	0.456 [0.104]***	0.444 [0.105]***	0.413 [0.082]***	0.059 [0.040]	0.081 [0.034]**	0.074 [0.045]	8954
G. Britain and N. Ireland with N. Ireland Fixed Effect	0.491 [0.042]***	0.475 [0.044]***	0.485 [0.042]***	0.02 [0.015]	0.065 [0.013]***	0.065 [0.013]***	82908
Birth Cohort Polynomial Controls	Quartic	Quartic	Quartic	Quartic	Quartic	Quartic	
Age Polynomial Controls	None	Quartic	None	No	Quartic	None	
Age Dummies	No	No	Yes	No	No	Yes	

 Table 1

 Estimated Effect of Minimum School Leaving Age on Age Finished Full Time Education and Log Annual Earnings Great Britain and Northern Ireland, Ages 25 - 64, 1935 - 1965

Notes: The dependent variables are age left full-time education and log annual earnings. Each regressions includes controls for a birth cohort quartic polynomial and and indicator whether a cohort faced a school leaving age of 15 at age 14. Columns (2), (3), (5), and (6) also include age controls: a quartic polynomial and fixed effects where indicated. Each regression includes the sample of 25 to 64 year olds from the 1984 through 2006 General Household Surveys, who were aged 14 between 1935 and 1965. Data are first aggregated into cell means and weighted by cell size. Regressions are clustered by birth cohort and region (Britian or N. Ireland).

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	I	Returns to Schooling: OLS	5	Return	ns to Compulsory School	ng: IV	Initial Sample Size
Great Britain	0.073 [0.001]***	0.075 [0.001]***	0.075 [0.001]***	0.112 [0.034]***	0.111 [0.033]***	0.125 [0.040]***	73954
Norther Ireland	0.101 [0.003]***	0.105 [0.003]***	0.105 [0.003]***	0.129 [0.076]*	0.18 [0.062]***	0.179 [0.096]*	8954
G. Britain and N. Ireland with N. Ireland Fixed Effect	0.081 [0.002]***	0.085 [0.002]***	0.085 [0.002]***	0.041 [0.032]	0.133 [0.027]***	0.135 [0.028]***	82908
Birth Cohort Polynomial Controls	Quartic	Quartic	Quartic	Quartic	Quartic	Quartic	
Age Polynomial Controls	None	Quartic	None	No	Quartic	None	
Age Dummies	No	No	Yes	No	No	Yes	

 Table 2

 OLS and IV Returns to (Compulsory) Schooling Estimates for Log Annual Earnings

 Great Britain and Northern Ireland, Ages 25 - 64, 1935 - 1965

Notes: The dependent variable is log annual earnings. Each regressions includes controls for a birth cohort quartic polynomial and age left full time education (instrumented by an indicator whether a cohort faced a school leaving age of 15 at age 14 in Columns (4) through (6)). Columns (2), (3), (5), and (6) also include age controls: a quartic polynomial and fixed effects where indicated. Each regression includes the sample of 25 to 64 year olds from the 1983 through 2006 General Household Surveys who were aged 14 between 1935 and 1965. Data are first aggregated into cell means and weighted by cell size. Regressions are clustered by birth cohort and region (Britian or N. Ireland).

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Table 3 First Stage Effects of Compulsory Schooling on Education Attainment and Earnings for the U.S., Canada, and the U.K.

	(1)	(2)	(3)	(4)	(5)	(6)		
	1st Stage	Effects of Dropout	Ages on Schooling	Reduced	Form Coefficients	s on Earnings		
	Full Sample	Sample with < High School	Sample with > High School	Full Sample	Sample with < High School	Sample with > High School		
	τ	United States [190	1 - 1961 Birth Cohor	ts aged 25-64 in the	1950 - 2000 Cens	uses]		
Dependent Variable	Num	ber of Years of Scho	ooling		Log Weekly Wage	2		
Minimum School Leaving Age at age 14	0.110 [0.0070]***	0.100 [0.0097]***	0.003 [0.0027]	0.016 [0.0015]***	0.010 [0.0024]***	0.003 [0.0017]*		
Initial Sample Size	2,814,203	727,789	1,173,880					
F-test: Schl. leaving age coeff. is zero	243.5							
		Canada [1911 -	1961 Birth Cohorts a	nged 25-64 in the 19'	71 - 2001 Censuse	es]		
Dependent Variable	Num	ber of Years of Scho	ooling		Log Annual Wage	2		
Minimum School Leaving Age at age 14	0.130 [0.0154]***	0.130 [0.0129]***	-0.026 [0.0114]**	0.012 [0.0037]***	0.012 [0.0047]**	-0.003 [0.0049]		
Initial Sample Size	854,243	355,299	298,342					
F-test: Schl. leaving age coeff. is zero	70.5							
	τ	nited Kingdom	1921 - 1951 Birth Col	orts aged 28-64 in t	he 1984 - 2006 G	HHS]		
Dependent Variable		Left Full-Time Educ		0	Log Annual Wage	-		
Minimum School Leaving Age	0.489	0.627	0.020	0.053	0.049	0.012		
at age 14	[0.049]***	[0.037]***	[0.091]	[0.017]***	[0.028]*	[0.027]		
Initial Sample Size	82,908	45,359	37,549					
F-test: Schl. leaving age coeff. is zero	101.2							
		Britian [1921	- 1951 Birth Cohorts	aged 28-64 in the 19	984 - 2006 GHHS]		
Dependent Variable	Age	Left Full-Time Educ	cation		Log Annual Wage	-0.003 [0.0049] HHS] 0.012 [0.027]		
Minimum School Leaving Age at age 14	0.436 [0.064]***	0.553 [0.107]***	-0.151 [0.198]	0.047 [0.018]**	0.060 [0.016]***			
Initial Sample Size	73,954	40,692	33,262					
F-test: Schl. leaving age coeff. is zero	46.5							

Note: Regressions in the top three panels include fixed effects for birth year, region (state, province, Britain/N.Ireland), survey year, sex, and a quartic in age. The U.S. results also include a dummy variable for race, and state controls for fractions living in urban areas, black, in the labor force, in the manufacturing sector, female, and average age based on when a birth cohort was age 14. Provincial controls for Canada include fractions in urban areas, in the manufacturing sector, and controls for pre capital public and school expenditures. Data are grouped into means by birth year, nation, sex, race (for the U.S.) and survey year and weighted by cell population size. Huber-White standard errors are shown from clustering by region and birth cohort. Single, double, and triple asterizes indicate significant coefficients at the 10 percent, and 1 percent levels respectively. The omitted variable indicates ability to drop out at age 13 or lower for the U.S. and Canada, and 14 or less for the U.K. Samples include aldults aged 28 to 64, horn between 1921 and 1951. Dependent variable in Column 3 for Canada is 1 = some post secondary schooling, 0 otherwise. The last panel shows results with only the British sample, using a quartic birth cohort polynomial instead of cohort fixed effects.

	(1)	(2)	(3)	(4)
	OLS Full Sample	IV with Regional Controls	IV with Regional Trends	IV with Regional Trends and Regional Controls
Dependent Variable	United States in			
Log Weekly Earnings (all workers)	0.078	0.142	0.175	0.405
Log weekly Lannings (an workers)	[0.0005]***	[0.0119]***	[0.0426]***	[0.7380]
Log Weekly Earnings (males)	0.070	0.127	0.074	0.235
	[0.0004]***	[0.0145]***	[0.0384]*	[0.1730]
Log Weekly Earnings (black males)	0.074	0.172	0.119	0.264
	[0.0004]***	[0.0137]***	[0.0306]***	[0.1295]**
		011-1961 Birth Cohort the 1971-2001 Census		
Log Annual Earnings (all workers)	0.099	0.096	0.095	0.142
	[0.0007]***	[0.0254]***	[0.1201]	[0.0652]**
Log Annual Earnings (males)	0.087	0.124	-0.383	0.115
	[0.0008]***	[0.0284]***	[1.1679]	[0.0602]*
		n [1921-1951 Birth Co 1 the 1979 - 2006 GHH		
Log Annual Earnings (all workers)	0.085	0.108	-0.056	NA
	[0.002]***	[0.0328]***	[0.0468]	
Log Annual Earnings (males)	0.065	0.053	-0.032	NA
	[0.0021]***	[0.039]	[0.0475]	
		21-1951 Birth Cohorts 1 the 1979 - 2006 GHH	-	
	OLS	RD		
Log Annual Earnings (all workers)	0.083 [0.003]***	0.101 [0.0421]**	NA	NA
Log Annual Earnings (males)	0.063 [0.0021]***	0.110 [0.0551]*	NA	NA

Table 4 OLS, IV-DD, and IV-RD Estimates of the Returns to (Compulsory) Schooling for the U.S., Canada, and the U.K.

Note: Regressions in the top three panels include fixed effects for birth year, region (state, province, Britain/N.Ireland), survey year, sex, and a quartic in age. The U.S. results also include a dummy variable for race, and state controls for fractions living in urban areas, black, in the labor force, in the manufacturing sector, female, and average age based on when a birth cohort was age 14. Provincial controls for Canada include fraction in urban areas, in the manufacturing sector, and controls for per capita public and school expenditures. Data are grouped into means by birth year, nation, sex, race (for the U.S.) and survey year and weighted by cell population size. Huber-White standard errors are shown from clustering by region and birth cohort. Single, double, and triple asterixes indicate significant coefficients at the 10 percent, 5 percent, and 1 percent levels respectively. The omitted variable indicates ability to drop out at age 13 or lower for the U.S. and Canada, and 14 or less for the U.K. The last panel repeats regression discontinuity results from Table 2 using the British sample only and a quartic birth cohort polynomial instead of cohort fixed effects.

	(1)	(2)	(3)
	Mean <hs sample<="" th=""><th>OLS</th><th>IV Full Sample</th></hs>	OLS	IV Full Sample
Country (Schooling Variable)		Health Outcomes	
United States (Total Years of Schooing)			
Physical or Mental Health Disability That Limits Personal Care	0.092	-0.014 [0.0003]***	-0.025 [0.0058]***
Disability That Limits Mobility	0.128	-0.020 [0.0004]***	-0.043 [0.0070]***
United Kingdom (Age Left Full Time Education)			
Self Reported Poor Health	0.150	-0.037 [0.0016]***	0.007 [0.0084]
Self Reported Good Health	0.564	0.065 [0.0021]***	-0.010 [0.0114]
	Other	Socialeconomic Ou	tcomes
United States (Schooling Variable: Total Years of Section 2015)	chooing)		
Unemployed	0.064	-0.004 [0.0002]***	-0.005 [0.0040]
Receiving Welfare	0.067	-0.013 [0.0002]***	-0.011 [0.0024]***
Below Poverty Line	0.220	-0.023 [0.0002]***	-0.064 [0.0085]***
Canada (Total Years of Schooling)			
Unemployed; looking for work	0.062	-0.038 [0.0044]***	-0.010 [0.003]***
Below Low-Income Cut-off	0.227	-0.038 [0.0004]***	-0.026 [0.0038]***
Below Low-Income Cut-off United Kingdom (Age Left Full Time Education)	0.227		

 Table 5

 OLS and IV Estimates for Effects of (Compulsory) Schooling on Socialeconomic Outcomes

Note: All Regressions include fixed effects for birth year, region (state, province, Britain/N.Ireland), survey year, sex, and a quartic in age. The U.S. results also include a dummy variable for race, and state controls for fractions living in urban areas, black, in the labor force, in the manufacturing sector, female, and average age based on when a birth cohort was age 14. Provincial controls for Canada include fraction in urban areas, in the manufacturing sector, and controls for per capital public and school expenditures. Data are grouped into means by birth year, nation, sex, race (for the U.S.) and survey year and weighted by cell population size. Huber-White standard errors are shown from clustering by region and birth cohort. Single, double, and triple asterixes indicate significant coefficients at the 10 percent, 5 percent, and 1 percent levels respectively. See text for more data specifics.

0.066

-0.025

[0.0024]***

0.012

[0.0137]

Receiving Income Support

	(1)	(2)	(3)	(4)	(5)
Discount Rate		Annual Retun		Baseline Forgone Earnings	Peak Income for Early Dropout
			United States		
	0.154	0.11	0.07		(at Age 54)
0.03	103,593	72,485	45,200	7,525	34,243
0.05	68,472	47,911	29,876	7,525	34,243
0.07	48,236	33,751	21,046	7,525	34,243
0.295	7,525				
			Canada		
	0.129	0.11	0.07		(at Age 47)
0.03	82,572	69,616	43,411	7,525	30,827
0.05	55,029	46,395	28,931	7,525	30,827
0.07	38,972	32,857	20,489	7,525	30,827
0.252	7,525				
			United Kingdo	m	
	0.134	0.11	0.07		(at Age 56)
0.03	110,299	87,095	54,310	7,525	43,543
0.05	72,435	57,197	35,667	7,525	43,543
0.07	50,660	40,003	24,945	7,525	43,543
0.295	7,525				
	* estimated ann	ual return			

Table 6Average Financial Gain from Dropping Out One Year Later,
Measured in Present Value (2000 U.S. dollars)

Notes: Projected wage profiles among 15 and 16 year-old dropouts between the ages of 16 and 64 are shown in Figure 5. Column 1 shows the estimated annual returns to compulsory schooling from the instrumental variable regressions used to create these profiles. Column 1 also converts the annual profile differences to present value at age 15. Columns 2 and 3 show present value gains assuming alternative annual returns to compulsory schooling (.11 and .07 rspectively). The baseline wage for a 15 year old dropout at age 15 is \$7,525. The fourth row for each country shows the discount factor needed to generate present value gains equal to this amount. Column 5 displys the projected peak earnings for a 15-year-old dropout.